

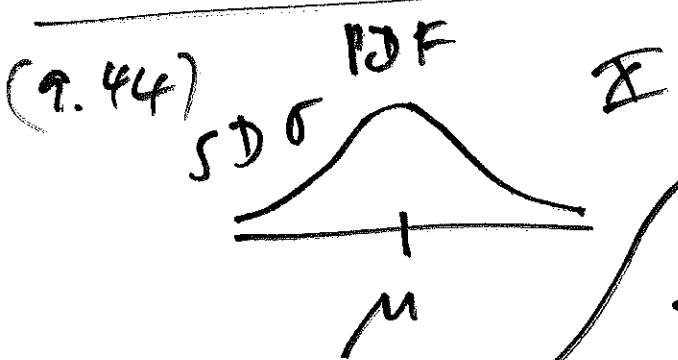
This pairson, Negative Binomial,
time: Geometrical, Normal, Gamma,
next
time: Beta, Multinomial, Bivariate Normal

①

THAT 2 due **TONIGHT**; Quiz 7 due **tomorrow night.**

Quiz 8 due next Tue

today: extra notes
p. 243+



de Moivre (1710)

$$f_X(x | \mu, \sigma^2) =$$

$$\frac{1}{\sigma\sqrt{2\pi}} \exp\left[-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2\right]$$

↑
converting
to standard
units
(Z)

"Normal" 1890
(Karl Pearson)

"Gaussian" (1790)

$$-\infty < \mu < +\infty$$

$$\sigma \geq 0$$

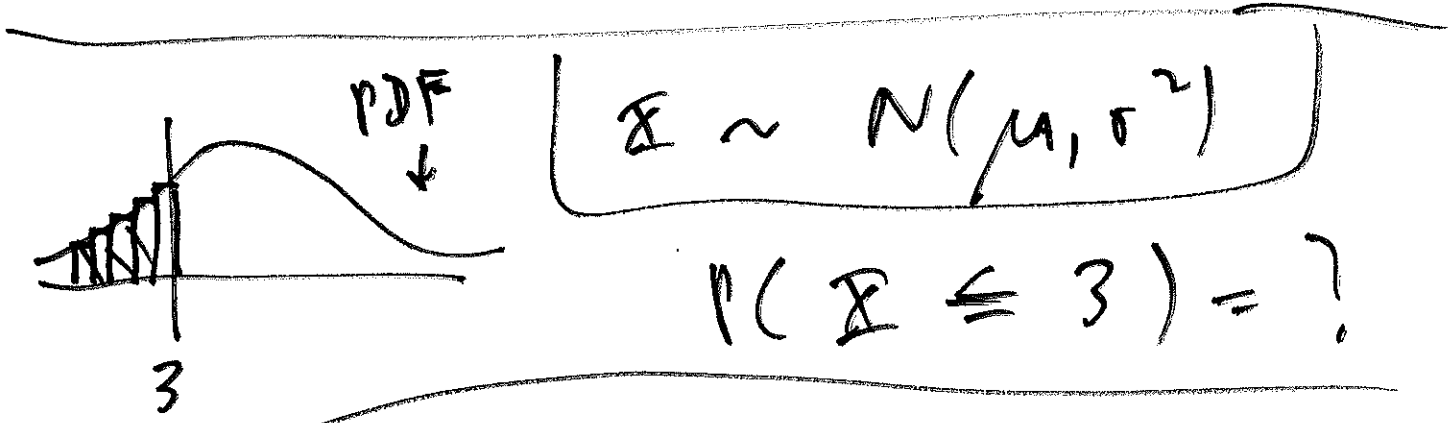
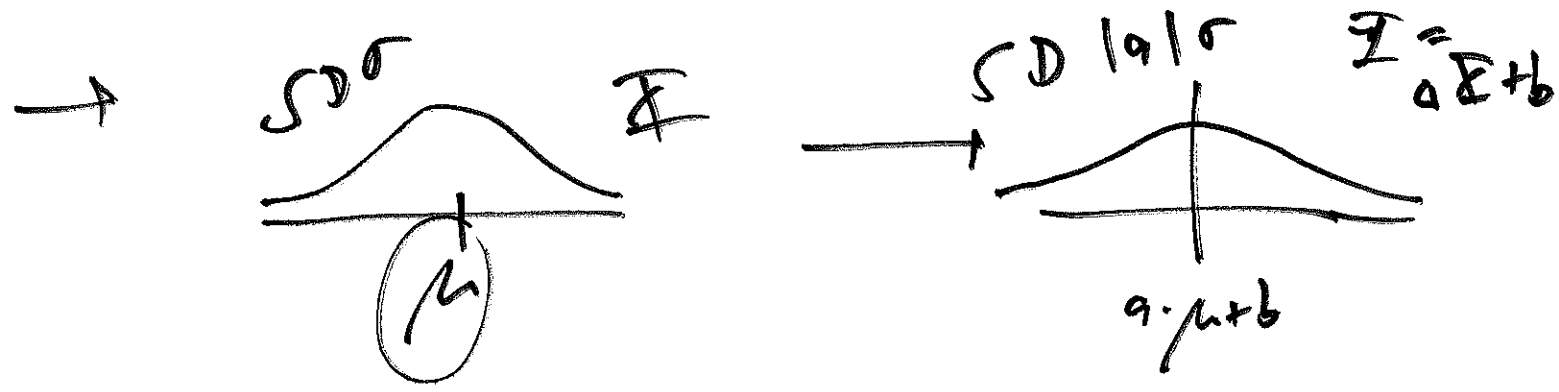
$$\sigma = 0$$



let's use
 $\sigma > 0$

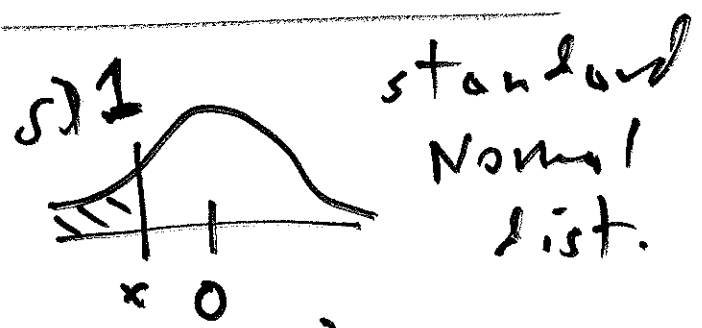
$(X | \mu, \sigma^2) \sim N(\mu, \sigma^2),$

$Z = aX + b, (a \neq 0)$ fixed constants



CDF of Normal (μ, σ^2) has no closed form

$\Phi(x) = \int \dots$



$\Phi(x) = \int_{-\infty}^x \frac{1}{\sqrt{2\pi}} \exp\left[-\frac{t^2}{2}\right] dt$

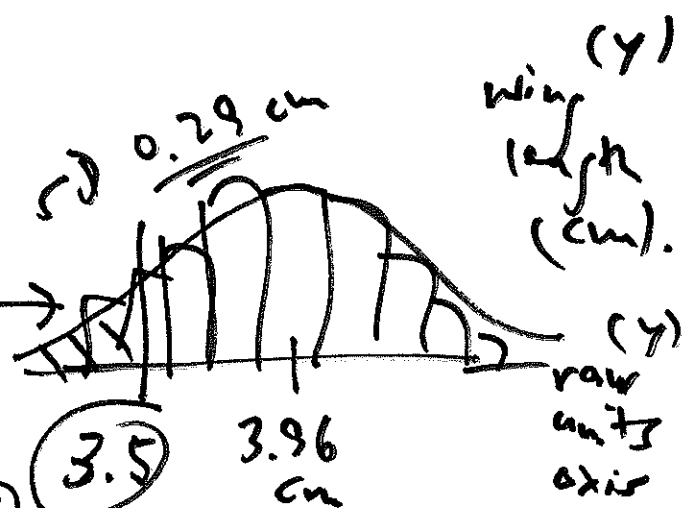
$$x = \begin{cases} 1 & \text{if } y \leq 3.5 \\ 0 & \text{else} \end{cases}$$

- $y_1 = 4.1$
- $y_2 = 3.3$
- \vdots
- $y_n = 4.7$

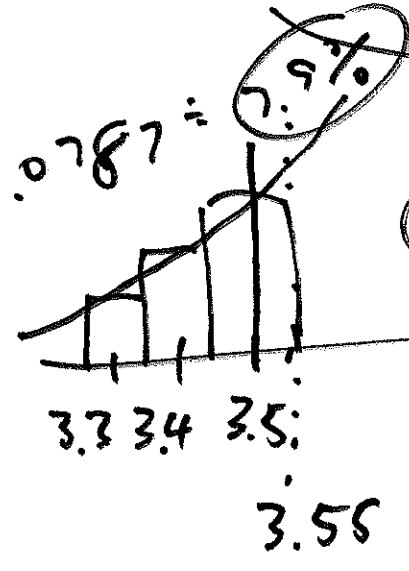
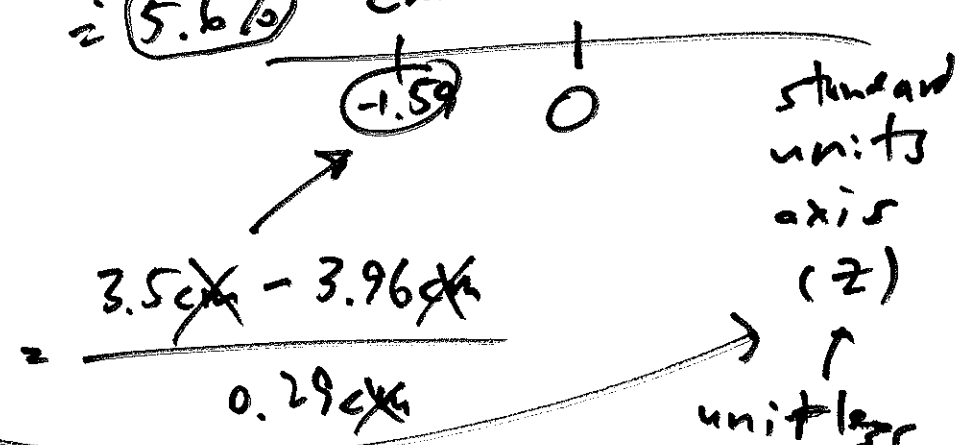
- 0
- 1
- 1
- 1
- 0

$n = 103$

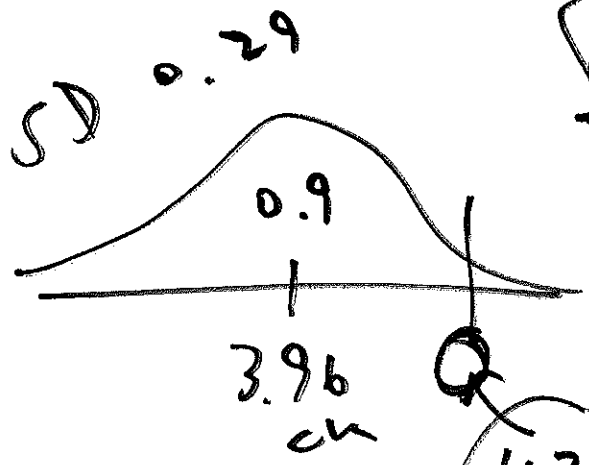
mean $\bar{y} = 3.96 \text{ cm}$
 SD $s = 0.29 \text{ cm}$



$$z = \frac{y - \bar{y}}{s}$$



continuity correction

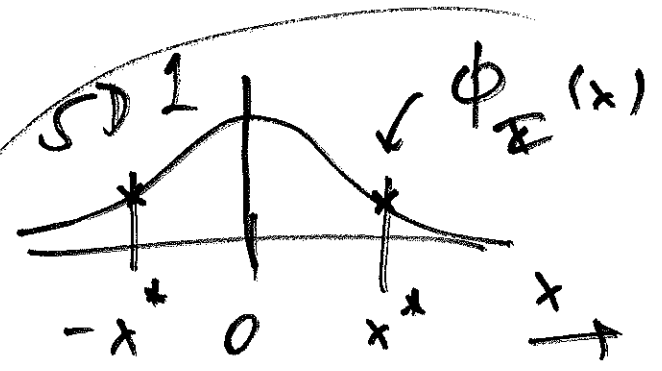


4.33 cm

Q: what is the 90th percentile of the wire length dist.?
inverse CDF

$$\Phi(x) = \int_{-\infty}^x \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{t^2}{2}\right) dt$$

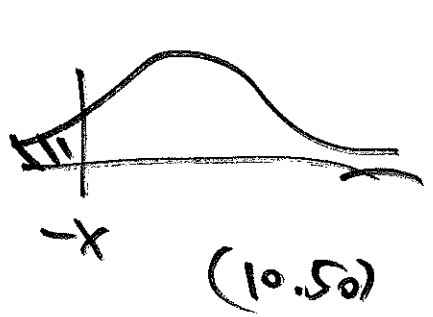
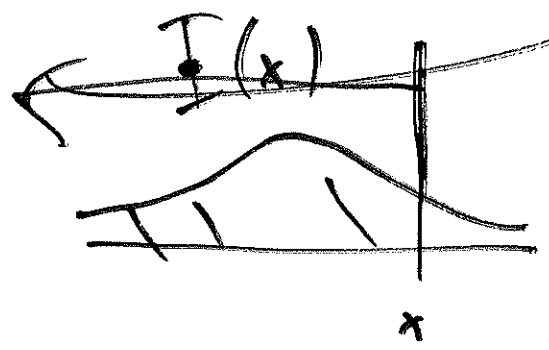
$$\Phi^{-1}(p) = ?$$



for all $x \geq 0$

$$\Phi(x) = \Phi(-x)$$

$$\Phi(x) + \Phi(-x) = 1$$



$$= 1$$

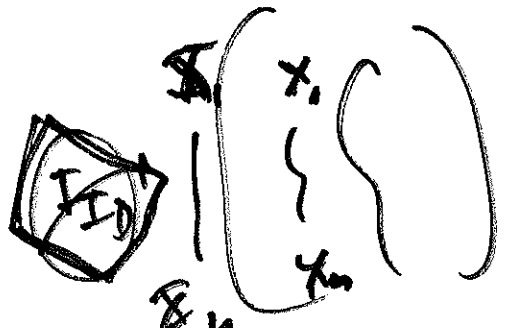
for all x

μ

sample

intuition:

\bar{X}_n is a good estimate of μ



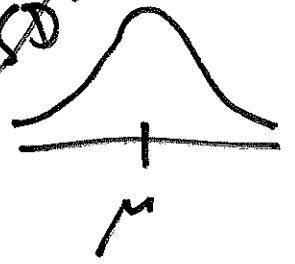
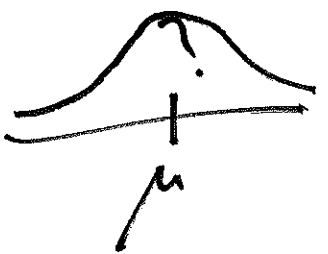
mean $\mu = ?$
SD $\sigma = ? < \infty$

SD $s =$

sample hist.

$\sigma = SE = ?$

PDF of \bar{X}_n



$$E(\bar{X}_n) = E\left(\frac{1}{n} \sum_{i=1}^n X_i\right) = \frac{1}{n} E\left(\sum_{i=1}^n X_i\right)$$

$$= \frac{1}{n} \sum_{i=1}^n E(X_i) \stackrel{(2)}{=} \frac{1}{n} \sum_{i=1}^n \mu = \mu \checkmark$$

$E(\bar{X}_n) = \mu \iff \bar{X}_n$ is unbiased for μ

$$SD(\bar{X}_n) = \sqrt{E(\bar{X}_n)} = \frac{\sigma}{\sqrt{n}} = ? \quad (6)$$

$$SD\left(\frac{1}{n} \sum_{i=1}^n X_i\right) = \frac{1}{n} SD\left(\sum_{i=1}^n X_i\right) = ?$$

$$V(\bar{X}_n) = V\left(\frac{1}{n} \sum_{i=1}^n X_i\right)$$

$$= \frac{1}{n^2} V\left(\sum_{i=1}^n X_i\right)$$

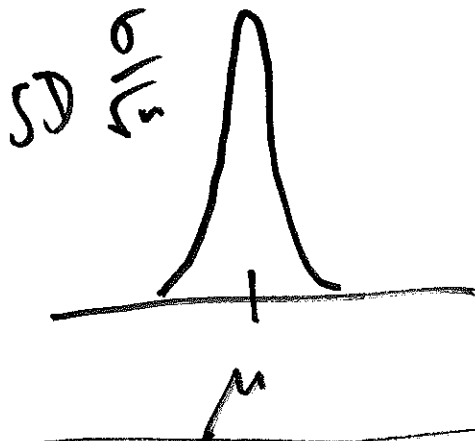
$$\stackrel{(1)}{=} \frac{1}{n^2} \sum_{i=1}^n V(X_i)$$

$$\stackrel{(2)}{=} \frac{1}{n^2} \sum_{i=1}^n \sigma^2 = \frac{n\sigma^2}{n^2} = \frac{\sigma^2}{n}$$

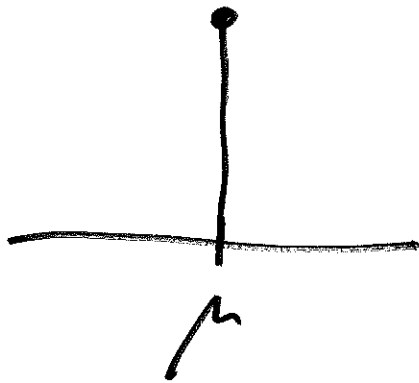
$$SD(\bar{X}_n) = \sqrt{E(\bar{X}_n)} = \sqrt{V(\bar{X}_n)}$$
$$= \sqrt{\frac{\sigma^2}{n}} = \frac{\sigma}{\sqrt{n}} \quad \leftarrow \begin{array}{l} \text{square} \\ \text{root} \\ \text{law} \end{array}$$



PDF of \bar{X}_1 ($\mu = \mu$, PDF)



PDF of \bar{X}_n ($n > 1$)



$n \rightarrow \infty$

assumption:
 $\sigma < \infty$