

08/30/19

AMS 131

Lecture 15

The set of values a Markov chain can take on is called its state space S , finite or infinite
(Can also have Markov chains unfolding in continuous time, ex: X_t = stock price at time t
 $t = \text{time}$)

1. (X_1, X_2, \dots) finite Markov chain \rightarrow

$$\begin{aligned} P(X_1 = x_1, \dots, X_n = x_n) &= P(X_1 = x_1) \cdot P(X_2 = x_2 | X_1 = x_1) \\ &\quad \cdot P(X_3 = x_3 | X_2 = x_2) \dots \\ &\quad \cdot P(X_n = x_n | X_{n-1} = x_{n-1}) \end{aligned}$$

Suppose you have a finite Markov chain with k possible states numbered $1, \dots, k$

$(k \text{ integer } \geq 2) \rightarrow \{P(X_{n+1} = j | X_n = i) \mid i = 1, \dots, k, n = 1, 2, \dots\}$
is called the transition distribution of the M.C.

If $P(X_{n+1} = j | X_n = i)$ is the same for all n ,
the transition distribution is said to be stationary (time-homogeneous)

\hookrightarrow if so, $P_{ij} \triangleq P(X_{n+1} = j | X_n = i)$ completely characterize the Markov chain's behavior

$$P = \begin{matrix} & \text{to state} \\ \text{from state} & \left[\begin{matrix} P_{11} & P_{12} & \dots & P_{1K} \\ \vdots & \ddots & \dots & \\ P_{K1} & P_{K2} & \dots & P_{KK} \end{matrix} \right] \end{matrix} \begin{matrix} \text{transition} \\ \text{matrix } P \end{matrix}$$

Row must add up to 1 \rightarrow chain goes somewhere
all elements $0 < P_{ij} < 1$ (probabilities)

A square matrix $P_{k \times k}$ with non-negative entries and all row sums = 1 is a stochastic matrix.

Gene inheritance is markovian, the genetic makeup for you is only determined by your parents, ancestors are irrelevant because they essential make up your parents.

Say a gene of interest has alleles A and a,

then a state in the Markov chain is in the form:

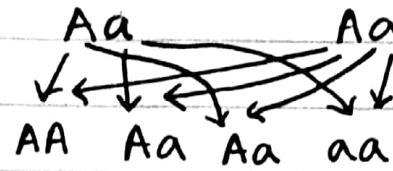
{ allele 1 from allele 2 from allele 1 from allele 2 from }
 parent 1 parent 1, parent 2 parent 2
 ex: {Aa, Aa}

ignoring order (not important in inheritance): 6 states
 {AA, AA}, {AA, Aa}, {AA, aa}, {Aa, Aa}, {Aa, aa}, {aa, aa}

one possible

inheritance

sequence



Offspring gets A or a from each parent (independent),
 each with probability $\frac{1}{2}$

transition matrix

from to {AA, AA} {AA, Aa} {AA, aa} {Aa, Aa} {Aa, aa} {aa, aa}

{AA, AA}	1	0	0	0	0	0
{AA, Aa}	$\frac{1}{4}$	$\frac{1}{2}$	0	$\frac{1}{4}$	0	0
{AA, aa}	0	0	0	1	0	0
{Aa, Aa}	$\frac{1}{16}$	$\frac{1}{4}$	$\frac{1}{8}$	$\frac{1}{4}$	$\frac{1}{4}$	$\frac{1}{16}$
{Aa, aa}	0	0	0	$\frac{1}{4}$	$\frac{1}{2}$	$\frac{1}{4}$
{aa, aa}	0	0	0	0	0	1

Random walk: you're watching a particle move on the integers $S = \{\dots -2, -1, 0, 1, 2, \dots\}$ over time

• wherever it is at time n it moves

left 1 unit with prob p_1

$$0 < p_1 < 1$$

stays with prob p_2

$$\sum_{i=1}^3 p_i = 1$$

right 1 unit with prob p_3

this is a markov chain because its future position only depends on where it is, not its previous

Lecture 15 (cont.)

from	to	...	-2	-1	0	1	2	...	transitions
...	matrix
-2	~	P ₂	P ₃	0	0	0	0	~	
-1	~	P ₁	P ₂	P ₃	0	0	0	~	
0	~	0	P ₁	P ₂	P ₃	0	~	= P	
1	~	0	0	P ₁	P ₂	P ₃	0	~	
2	~	0	0	0	P ₁	P ₂	0	~	
...	

band matrix: only nonzero entries are on the main diagonal and one diagonal away from the main

- 3 diagonals: tri-diagonal

all the main diagonal entries are the same P_2
one above: P_3 , one below: P_1

Start the random walk, at 0 let it go, where is it likely to be at time n , n large

Suppose: $(P_1, P_2, P_3) = (0.1, 0.3, 0.6)$

will eventually drift to $+\infty$

$(0.5, 0.25, 0.25)$ drift to $-\infty$

$(\frac{1}{3}, \frac{1}{3}, \frac{1}{3}) = ?$ as $n \rightarrow \infty$, each integer is visited infinitely many times and the expected time for it to return to 0 is also infinite \rightarrow "too much freedom"

We can restrict the bounds to: for $k \geq 1$

$$S = \{-k, -k+1, \dots, -1, 0, 1, \dots, k-1, k\}$$

at the boundary $\begin{array}{ccc} \text{---} & \text{---} & \text{---} \\ k-1 & k & k+1 \end{array}$ not possible

so we can wrap around: if you move to $(k+1)$, go to $-k$ and if move to $-(k+1)$, go to k

transition matrix, $K=2$

$$\begin{matrix} -2 & -1 & 0 & 1 & 2 \end{matrix} \quad \text{to}$$

$$\begin{matrix} -2 & p_2 & p_3 & 0 & 0 & p_1 \end{matrix}$$

$$\begin{matrix} -1 & p_1 & p_2 & p_3 & 0 & 0 \end{matrix}$$

$$\begin{matrix} 0 & 0 & p_1 & p_2 & p_3 & 0 \end{matrix}$$

$$\begin{matrix} 1 & 0 & 0 & p_1 & p_2 & p_3 \end{matrix}$$

$$\begin{matrix} 2 & p_3 & 0 & 0 & p_1 & p_2 \end{matrix}$$

another idea is

to make the

boundaries

impossible to

cross, hard stop.

$$P_{ij}^{(m)} = P(\text{chain moves from } i \text{ to } j \text{ in } m \text{ steps}) \\ = P(X_{n+m} = j \mid X_n = i)$$

finite Markov chain with stationary transition distributions and transition matrix $P \rightarrow P_{ij}^{(m)}$

is the (i, j) entry of the matrix $P_i^{(m)}$, which is called the m -step transition matrix.

$\{\text{AA}, \text{AA}\}$ and $\{\text{aa}, \text{aa}\}$ has the property that once the chain is in that state, it cannot go anywhere else — state i when $p_{ii} = 1$ any state with $p_{ii} = 1$, is called an absorbing state

In the genetic chain, 1, 2, 4 have positive probability of moving to state 1 in 2 steps, and the same for moving to state 6 in 2 steps.

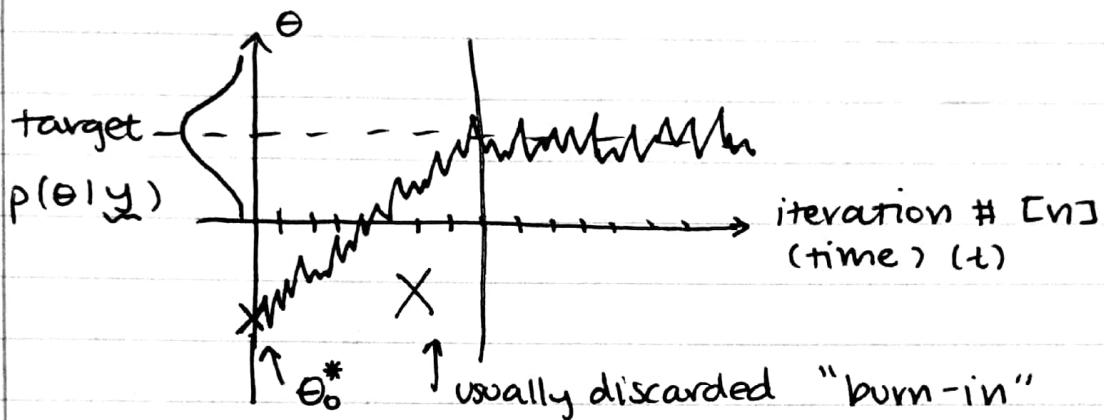
Markov chains that settle down to a single, stable, long run distribution (equilibrium distribution) are called steady state (stationary dist). ← confusing with time-homogeneous

Where should the Markov chain start?

- either initialize it at a deterministic value or make a random draw from the initial distribution

Lecture 15 (cont.)

Any vector \mathbf{x} of non-negative numbers that add up to 1 is called a probability vector - any such vector whose components specify that a Markov chain will be in each possible state at time 1 is the initial distribution of the chain.



After 1 time step (iteration) the probability dist. over the Markov chain's possible states is \mathbf{x} , after 2 iterations, the chain's dist is: $\mathbf{x} \mathbf{P} \mathbf{x}$, after $(m+1)$, $\mathbf{x} \mathbf{P}^m$, looking for it to converge to a unique dist. as $m \rightarrow \infty$, resulting in its equilibrium distribution

$$\text{If we choose } \mathbf{x} \mathbf{P} = \mathbf{x}, \text{ then } \mathbf{x} \mathbf{P}^2 = (\mathbf{x} \mathbf{P}) \mathbf{P} \\ \lim_{m \rightarrow \infty} \mathbf{x} \mathbf{P}^m = \mathbf{x}$$

$$\mathbf{x} \mathbf{P}^3 = (\mathbf{x} \mathbf{P}^2) \mathbf{P} = \mathbf{x} \mathbf{P} = \mathbf{x}$$

Markov chain with transition matrix $\mathbf{P} \rightarrow$ only probability vector \mathbf{x} such that $\mathbf{x} \mathbf{P} = \mathbf{x}$ is an equilibrium dist. for the chain

- under additional conditions on \mathbf{P} , such an eq. dist. will be unique

$\sum_{k=1}^n p_k = 1$ eigen-analysis of P

$\sum_{k=1}^n x_k = \lambda x \leftarrow$ an eigenvector } both
 \leftarrow an eigenvalue "right"

↓ "left" eigenvector $v_P = \sum_{k=1}^n v_k \leftarrow$ eigenvalue of 1

Given a square matrix P , any vector $\sum_{k=1}^n v_k$ satisfying $\sum_{k=1}^n P_{k1} v_k = \lambda_1 v_1$ is called a right eigenvector of P

with eigen value λ_1 and any vector $\sum_{k=1}^n v_k$ satisfying $\sum_{k=1}^n P_{k1} v_k = \lambda_L v_L$ is called a left eigenvector with eigenvalue λ_L

So given a transition matrix P for a Markov chain, an equilibrium dist. can be found by computing the left eigenvector $\sum_{k=1}^n v_k$ where the eigenvalue is 1. notice if:

$$\sum_{k=1}^n P_{k1} = 1 \quad \text{then} \quad (\sum_{k=1}^n v_k)^T = \sum_{k=1}^n v_k^T \\ = P^T \sum_{k=1}^n v_k^T = \sum_{k=1}^n v_k^T$$

so we can eigendecompose P^T instead of P

$\begin{pmatrix} 1 \\ 0 \\ 0 \\ 0 \\ 0 \\ 0 \end{pmatrix}$ and $\begin{pmatrix} 1 \\ 0 \\ 0 \\ 0 \\ 0 \\ 0 \end{pmatrix}$ correspond to the absorbing states (of genetics problem), this suggests there's a family of eq. dist. of the form:

$$(p, 0, 0, 0, 0, 1-p)^T \text{ for } 0 \leq p \leq 1$$

If you can find an integer $m \geq 1$ such that every element of P^m is strictly positive, then $\lim_{n \rightarrow \infty} P^n$ is a matrix with all rows equal to the unique stationary dist. \sum and no matter what the initial dist. is, its dist after n steps converges to \sum as $n \rightarrow \infty$

Case study: stock options

Companies can give signing bonuses often in the form of stock options: an opportunity to buy N shares of the company one year from now at a (known) price, S

If you think the stock will rise = sell at profit.

X = price of stock 1 yr from now. \leftarrow

↳ for now, discrete with two values: $X_1 < S$ and $X_2 > S$ let: \rightarrow not known

$p = P(X=X_2)$ the prob that stock value will rise

I = value of option for one share at $\$S$

If ($X=X_1 < S$), option is worthless and $I=0$

If ($X=X_2 > S$) then the option is worth

$$(X_2 - S) \text{ thus } I = h(X) = \begin{cases} 0 & \text{if } X=X_1, \\ (X_2 - S) & X_2 \end{cases}$$

- Can compare stock option to investing in a bond that pays $\alpha\%$ / year \longrightarrow

• A fair measure of the worth of the option would be the present value of I , defined to be the number c such that $E(I) = (1+\alpha) \cdot c$

$$E(I) = 0 \cdot (1-p) + (X_2 - S)p = (X_2 - S)p$$

$$(1+\alpha) \cdot c = (X_2 - S)p \quad \text{comparing}$$

$$\text{need } \longrightarrow \quad c = \left(\frac{X_2 - S}{1+\alpha} \right)p \quad \begin{matrix} \text{stock options} \\ \text{to bonds} \end{matrix}$$

assume the present value of I is equal in expectation to the current value of the stock.

(buying 1 share) = (investing same amount into bond)
and hold 1 yr.

$$E(I) = (1+\alpha) \cdot S$$

$$\text{but } E(I) = p \cdot X_2 + (1+\alpha)X_1 \stackrel{\Delta}{=} (1+\alpha) \cdot S$$

$$p = \frac{X_1 - (1+\alpha) \cdot S}{X_2 - X_1} = \frac{(1+\alpha)S - X_1}{X_2 - X_1}$$

So, fair price C of an option to buy one share : $C = \left(\frac{X_2 - S}{1 + \alpha} \right) \left(\frac{(1+\alpha)S - X_1}{X_2 - X_1} \right)$

$$S = \$200 \quad \alpha = 0.04 \leftarrow \text{realistic}$$

$$X_1 = \$180 \quad \text{downside } \$20 (-10\%) \quad C = \$20.19$$

$$X_2 = \$260 \quad \text{upside } \$60 (+30\%) \quad \sim 10\% \text{ of current}$$

C = risk neutral price of the option

Under the assumptions we made, you could now sell the option today (if you had it) at a fair price = \$20 (options trading)

An investment that allows people to buy or sell an option on a security is called a derivative

case study: building a good portfolio.

You have $\$I$ to invest to build a portfolio, optimal collection for companies A, C + bonds B

This depends on how much A, C will change

R_A = rate of return (change in price) for comp. A

R_C = " "

your portfolio: $(S_A \text{ shares of A stock } \$S_B \text{ worth})$
 $(S_C \text{ shares of C stock of bonds})$

1 year return: $S_A R_A + S_C R_C + \$r S_B$

fixed random fixed

$$E(R_A) = \mu_A \quad V(R_A) = \sigma_A^2 \quad (S_A, S_C, S_B) : \text{total}$$

$$E(R_C) = \mu_C \quad V(R_C) = \sigma_C^2 \quad \text{current value} = \$I$$

P_A = current price of A, $P_C = " "$ of C

$$\text{constraint: } S_A P_A + S_C P_C + S_B = I$$

$$\text{Pretend } R_A, R_C. \quad E(S_A R_A + S_C R_C + r S_B) = S_A \mu_A + S_C \mu_C + r S_B$$

$$\text{independent} \quad V(S_A R_A + S_C R_C + r S_B) = S_A^2 \sigma_A^2 + S_C^2 \sigma_C^2$$

\hookrightarrow we want: expected high, variance low (rare)

In practice need to sacrifice some expected return to keep downside risk small.

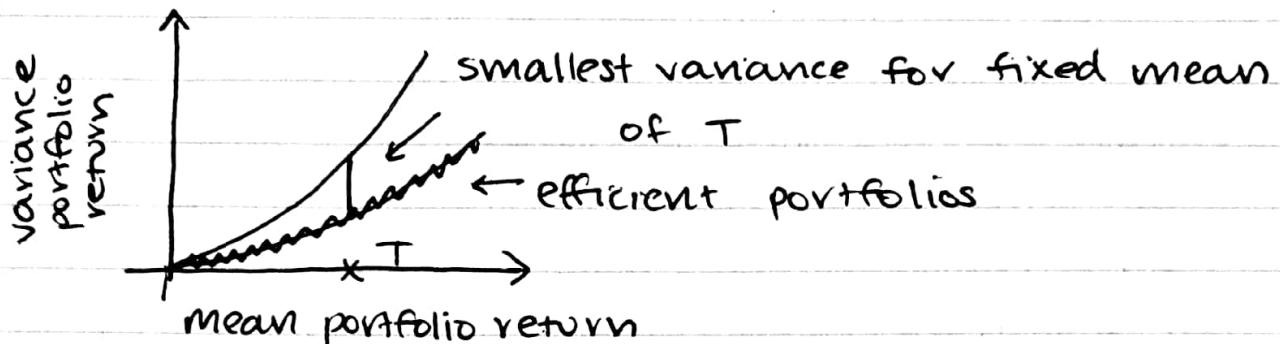
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Lecture 15 (cont.)

How should you optimally make this trade off?

optimize over (U, V)

1. optimize $pU + (1-p)V$ ✓ or vice versa
2. optimize on V subject to constraint on U



Suppose you set target $\$T$ for your expected return

$E(\text{return}) = T$ with smallest variance = best

↳ minimal variance portfolios are efficient

for fixed: $P_A, P_C, T, I, r, M_A, M_C, \sigma_A^2, \sigma_C^2$
find (S_A, S_C, S_B) to need to speculate

minimize $V = S_A^2 \sigma_A^2 + S_C^2 \sigma_C^2$ subject to

$$E = S_A M_A + S_C M_C + r S_B = T \text{ and}$$

$$S_A P_A + S_C P_C + S_B = I$$

instead of first derivatives, use Lagrange multipliers

$$I = \$100,000 \quad T = \$7,000 \quad (7\% \text{ reasonable return})$$

$$V = 2.55 \times 10^7 \$^2, \quad SD = \$5050$$

If dist. of return follows

bell curve, only 8% chance
of losing money.

